Highly Adaptive Lasso Mark

Motivation

Overview

Implementation

Application of HAL

Super Learning Propensity Score Estimation in TML

Efficient Plug-in Estimation

Nonparametric Bootstrap

Concluding Remarks Acknowledgements: Rachael Phillips and Lars van der Laan

Highly Adaptive Lasso In Causal Inference

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October 25, 2021, BASS-meeting

### Traditional Lasso Estimator

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Concluding Remarks • Advantages:

- L<sub>1</sub>-regularization performs both variable selection and penalized regression
- Interpretable
- Cross-validated selection of *L*<sub>1</sub>-norm/penalty parameter
- Disadvantages:
  - Reliance on parametric forms places strong assumptions on the functional relationships between variables

#### HAL Advantages

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- Nonparametric Bootstrap
- Concluding Remarks

- First estimator that guarantees asymptotically efficient estimation of any pathwise differentiable estimand<sup>1</sup> (e.g., the average causal effect or treatment-specific survival), without enforcing strong smoothness conditions.
- Assumptions are exceedingly mild, and expected to hold in almost every practical application.
- Can be implemented with standard Lasso software.
- Converges to true function at rate  $n^{-1/3}(\log n)^{d/2}$ .
- Accommodates a variety of function space specifications.

<sup>&</sup>lt;sup>1</sup>An estimand that is a weakly differentiable functional of the density of the data, the case for most causal inference estimands under positivity.  $\neg$ 

# Highly Adaptive Lasso (HAL)

Highly Adaptive Lasso

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Concluding Remarks A maximum likelihood estimator over all, or subset of, cadlag functions with finite variation norm.

#### Key Ingredients

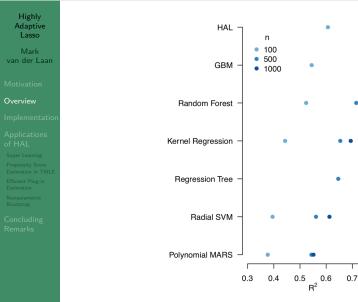
- Any stochastic relation/function we aim to learn from data can be approximated by linear combination (i.e., sum) of spline basis functions X → I(X > x<sub>i</sub>) for knot point x<sub>i</sub>.
- The variation norm (i.e., complexity) of a function is the  $L_1$ -norm.
- Optimize empirical performance over all such linear models under fixed *L*<sub>1</sub>-norm that is selected with cross-validation.

van der Laan, Mark. "A generally efficient targeted minimum loss based estimator based on the highly adaptive lasso." The International Journal of Biostatistics (2017).

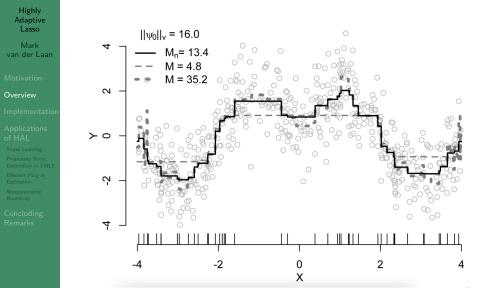
## Theoretically proven to approximate truth faster than known machine learning algorithms

...

0.8



#### Illustration in Low Dimensions



# Tuning HAL

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Concludin; Remarks HAL's computational cost is controlled by the number of basis functions, which can be as large as  $n * 2^{d-1}$ .

Options for constraining the functional form of the target function:

- Enforce a minimum proportion of 1's in basis function.
- Enforce a maximal order of interaction.
- Specify particular additive model.
- Enforce monotonicity for some of the functions.
- Enforce higher order splines and thereby smoothness.
- Discretize continuous covariates: fewer knot-points.
- Greedy screening: iteratively, selected j-th order basis functions generate j + 1-th order basis functions.

### Basic R hal9001 Functionality

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Concluding Remarks Load package and data: library(hal9001) data(mtcars)

Create numeric vector for dependent variable: Y <- mtcars[,"mpg"]</p>

③ Create dataframe or matrix of predictors: X <- mtcars[,c("cyl", "disp", "hp", "wt")]</pre>

Fit HAL:
 hal\_fit <- fit\_hal(X=X, Y=Y)
</pre>

*Note*: default max\_degree=3 considers no more than 3-way interactions, and default reduce\_basis=0 places no restrictions on the minimum proportion of 1's in basis functions.

#### Summary table of hal9001 HAL fit

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Concluding Remarks

#### summary(hal\_fit)\$table

coef <sup>‡</sup>	term $\diamond$		
35.4070	Intercept		
-4.0801	I(disp >= 440)		
-4.0118	I(disp >= 78.7)		
-2.7170	I(wt >= 1.513)		
-2.4454	I(wt >= 3.215)		
-1.8184	I(disp >= 71.1)		
-1.7208	I(hp >= 180)		
-1.6830	I(disp >= 95.1)		
-1.6039	I(hp >= 66)*I(wt >= 2.2)		
-1.5623	I(wt >= 2.2)		
1.3785	I(disp >= 351)		
-1.2444	I(hp >= 175)		
-1.1888	I(disp >= 301)		
-0.9026	I(hp >= 123)		
0.7336	I(hp >= 52)		
-0.5810	I(disp >= 120.1)*I(hp >= 97)		
-0.4395	I(disp >= 108)*I(hp >= 93)		

#### Specifying hal9001 model formulas

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Concluding Remarks Example: Observe  $O = (W_1, W_2, A, Y) \sim P_0$ 

R code: fit\_hal(Y, X, formula, data, ...)

#### Additive model formula:

$$ext{Y} \sim ext{. or } ext{Y} \sim ext{h(W1)} + ext{h(W2)} + ext{h(A)}$$

#### Bi-additive model formula:

 $Y \sim .^2$  or

 $\textbf{Y} \sim \texttt{h(W1)} + \texttt{h(W2)} + \texttt{h(A)} + \texttt{h(W1,W2)} + \texttt{h(W1,A)} + \texttt{h(W2,A)}$ 

#### Only interactions with A formula:

```
\mathtt{Y} \sim \mathtt{h}(.) + \mathtt{h}(.,\mathtt{A}) or
```

 $Y \sim h(W1)+h(W2)+h(A)+h(W1,A)+h(W2,A)$ 

 $\begin{array}{l} \textbf{Monotone} \uparrow \textbf{(i)} \downarrow \textbf{(d)} \text{ formula examples:} \\ \textbf{Y} \sim i(.) \text{ or } \textbf{Y} \sim i(.) + i(.,.) \text{ or } \textbf{Y} \sim i(\textbf{W1}) + d(\textbf{W2}) + i(\textbf{A}) \end{array}$ 

#### Possible HAL fits under various smoothness orders

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Concluding Remarks *Example*: Observe  $(W, A, Y) \sim P_0$ 

R code: fit\_hal(Y, X, data, formula, s, ...)

Example fits for 0-order smoothness, s=0:

Additive model:

$$Y = \mathbb{I}(W > 0.5) + \mathbb{I}(W > 0.3) + \mathbb{I}(A > 0)$$

Bi-additive model:

$$Y = \mathbb{I}(W > 0.5) + \mathbb{I}(A > 0) + \mathbb{I}(W > 0.5, A > 0)$$

#### Example fits for 1st-order smoothness, s=1: Additive model:

$$\begin{split} Y &= \mathbb{I}(W > 0.5)[W - 0.5] + \mathbb{I}(W > 0.3)[W - 0.3] + \mathbb{I}(A > 0)[A - 0] \\ \text{Bi-additive model:} \\ Y &= \mathbb{I}(W > 0.5)[W - 0.5] + \mathbb{I}(A > 0)[A - 0] + \mathbb{I}(W > 0.5, A > 0)[W - 0.5][A - 0] \end{split}$$

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#### Performance under Various Screening Options

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Concluding Remarks

- Binning: Reducing number of spline knot points, separately for 1-way, 2-way, 3-way basis functions.
- Greedy screening: Build interactions sequentially from screened basis functions of lower order interactions.

n = 2000; d = 12

Size of regression matrix  $(n \times p)$  up to 3-way interactions: Binning (100, 25, 5): p = 30,000No Binning: p = 600,000

Sequential (up to 2-way): first p=20,000 then p=2,000,  $R^2=0.88$ , MSE=0.376 No Sequential (up to 2-way): p=150,000,  $R^2=0.85$ , MSE=0.394

More basis functions does not imply better performance.

### Super Learner incorporating HAL

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Concluding Remarks

- By varying tuning parameters in HAL, one can include many HAL estimators in the library.
- The super learner will perform as well as the oracle choice among all these HAL estimators, and thereby achieves at minimal rate of convergence  $n^{-1/3}(\log n)^{d/2}$ .
- Can include other machine learning algorithms and parametric models as well.

## Meta-learning with HAL

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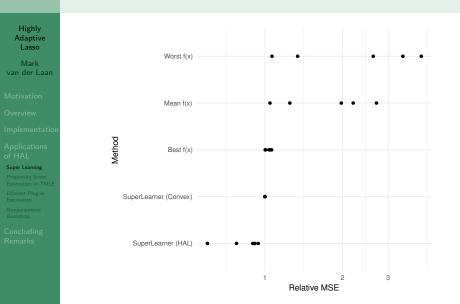
Concluding Remarks The super learner is defined by the the library of candidate estimators, loss function, and meta-learning algorithm.

#### Procedure for the Meta-HAL Super Learner

- **1** Perform meta-learning with HAL under specified  $L_1$ -norm.
- Obefine a discrete super learner that includes as candidates the L<sub>1</sub>-norm specific meta-HAL super learners, in order to optimally select the L<sub>1</sub>-norm of the HAL meta-learner (double cross-validation).

This implementation guarantees the final discrete-selected meta-HAL will perform as well as the optimally tuned meta-HAL super learner.

#### Meta-HAL Super Learner



# Outcome-regression weighted LASSO (OAL)

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Concluding Remarks Shortreed & Artefaie (2017) proposed outcome-regression weighted Lasso (OAL) for propensity score (PS) estimation:

• Fit unpenalized linear model for  $\mathbb{E}(Y|A, W)$ :

$$(\hat{\alpha}, \hat{\eta}) = \arg\min_{\alpha, \eta} I_n(\alpha; Y, A, W).$$

where  $\eta$  is the coefficient for A, and  $\alpha$  is the coefficient for W.

- Denote the coefficient for variable W<sub>j</sub> in the outcome regression with â<sub>j</sub>.
- Fit PS with Lasso using regularization term

 $\lambda \sum_j ||\alpha_j||^{-\gamma} ||\beta_j|| \text{ instead of usual } \sum_j | \beta_j |.$ 

## HAL-based OAL for PS Estimation

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Concluding Remarks The theoretical property of OAL relies on the *correct parametric formula*, which is often unknown in practice.

We extend OAL to outcome-regression weighted HAL (OHAL):

- Compute the outcome regression using Lasso, with dependent variable the outcome Y and features the basis functions \u03c6<sub>s,i</sub> and the treatment indicator A.
- **2** Get the outcome regression coefficients  $\alpha_{s,i}$  of  $\phi_{s,i}$ .
- Compute the propensity score using a Lasso logistic regression, with dependent variable A and features φ<sub>s,i</sub>. The L<sub>1</sub>-constraint for β<sub>s,i</sub>, the coefficient for φ<sub>s,i</sub>, is defined as the weighted L<sub>1</sub>-norm above.
- Tune the  $L_1$ -norm with C-TMLE.

# OHAL Performance based on Kang & Shafer (2007) Simulation

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Concludinį Remarks • Pre-treatment covariates  $(Z_{i1}, \dots, Z_{i4})$  are generated from uncorrelated standard normal distributions.

• Treatment indicator is then generated from a Bernoulli distribution with:

 $P(A_i = 1 | Z_i) = \exp(-Z_{i1} + 0.5Z_{i2} - 0.25Z_{i3} - 0.1Z_{i4})$ 

- Potential outcomes are generated by:  $Y_i^{(a)} = 210 + 27.4Z_{i1} + 13.7Z_{i2} + 13.7Z_{i3} + 13.7Z_{i4} + \epsilon$
- Thus, the value of the estimand, the ATE, is 0.

 $\epsilon \sim N(0,1)$ 

# Simulation to Assess OHAL Performance (Continued)

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Concluding Remarks • Only transformed covariates W are observed:

 $W_{i1} = \exp(Z_{i1}/2)$   $W_{i2} = z_{i2}/(1 + \exp(Z_{i1}) + 10)$   $W_{i3} = (Z_{i1}Z_{i3}/25 + 0.6)^3$  $W_{i4} = (Z_2 + Z_4 + 20)^2.$ 

• In our experiment, we also included an instrumental variable  $W_{i5}$ , and the treatment mechanism was modified to:

$$P(A_i = 1 | Z_i, W_i) = \exp(\frac{-Z_{i1} + 0.5Z_{i2} - 0.25Z_{i3} - 0.1Z_{i4}}{2} + W_{i5})$$

### **OHAL** Simulation Results

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Concluding Remarks We use main-term linear regression to create biased initial estimator  $\bar{Q}_n^0$  for TMLE/C-TMLE.

	TMLE-HAL	CTMLE-HAL	CTMLE-OHAL	Oracle
N=500	7.06	6.34	2.94	3.64
N=1000	4.42	3.55	1.40	1.70
N=2000	2.94	1.85	0.83	0.87

Table: MSE for each estimator across 200 replications with different sample size.

## HAL-TMLE is Asymptotically Efficient

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Concluding Remarks

- Due to the HAL-super learner converging to true nuisance parameter at faster rate than  $n^{-1/4}$ , the HAL-TMLE (using HAL for all nuisance parameters) is efficient in great generality (vdL, 15).
- The only necessary model assumptions are:
  - The true nuisance parameters have finite sectional variation norm
  - The loss functions of the true nuisance parameters are uniformly bounded, so that oracle inequality applies
  - The strong positivity assumption holds

# Example: Asymptotic efficiency of HAL-TMLE for treatment-specific mean / ATE

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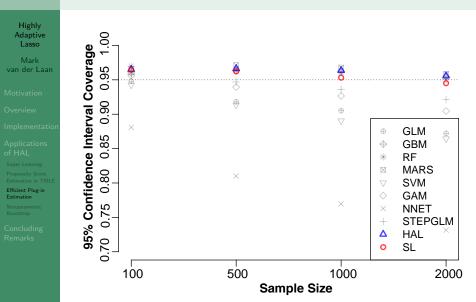
Nonparametric Bootstrap

Concluding Remarks Consider the HAL-TMLE of  $EY_1 = EE(Y | A = 1, W)$  based on  $(W, A, Y) \sim P_0$  in a nonparametric statistical model.

It is asymptotically efficient if

•  $\delta < P_0(A = 1 \mid W)$  for some  $\delta > 0$ 

- ②  $W \rightarrow E_0(Y \mid A = 1, W)$  and  $W \rightarrow P_0(A = 1 \mid W)$  are cadlag
- **③**  $W \rightarrow E_0(Y | A = 1, W)$  and  $W \rightarrow P_0(A = 1 | W)$  have finite sectional variation norm.



### Can we break HAL-TMLE?

# Undersmoothed HAL-MLE or Meta-HAL is efficient uniformly over large class of target estimands

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Concluding Remarks

- HAL-MLE is efficient for pathwise differentiable target estimands, if *L*<sub>1</sub>-norm is chosen large enough.
- Using HAL in the meta-learning step of the super learner to determine the best functional combination of candidate estimators: Meta-HAL SL.
- Meta HAL-SL is efficient for pathwise differentiable target estimands if *L*<sub>1</sub>-norm in meta-HAL is chosen large enough.
- Due to being an MLE, it solves a large class of score equations, in particular, efficient scores corresponding with target estimands. As a consequence, it can be analyzed as a plug-in TMLE, even though it is not targeted.

### Nonparametric Bootstrap of HAL-TMLE

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Concluding Remarks

- Fix *M* at the cross-validation selector *M<sub>n</sub>* or another selector (we propose a plateau selector!).
- Draw 10,000 samples of size *n* from empirical measure *P<sub>n</sub>*.
   For each bootstrap sample *P<sup>#</sup><sub>n</sub>*, recompute the HAL-TMLE(*M*), say **P**<sup>#\*</sup><sub>*n*,M</sub>.
- The HAL on bootstrap sample can be restricted to only include indicator basis functions that were selected by HAL-MLE(M) on original data.
- Use sampling distribution of ψ<sup>#\*</sup><sub>n,M</sub> = Ψ(**P**<sup>#\*</sup><sub>n,M</sub>), conditional on P<sub>n</sub>, to construct 0.95-confidence interval.
- Increase *M* till plateau in confidence interval for optimal coverage.

#### Bootstrap works for HAL-TMLE

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Concluding Remarks  Conditional on the data (P<sub>n</sub> : n ≥ 1), the bootstrap sampling distribution of ψ<sup>\*</sup><sub>n</sub> converges to optimal normal limit distribution N(0, σ<sup>2</sup><sub>0</sub>).

- The approximation error of bootstrap is driven by performance of nonparametric bootstrap for an empirical process indexed by Donsker class (i.e., cadlag functions with sectional variation norm bounded by *M*).
- This suggests robust finite sample behavior of the nonparametric bootstrap.

### Case Study

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Concluding Remarks Compare two confidence intervals for ATE  $EY_1 - EY_0$ :

- Wald-type
- HAL-TMLE bootstrap, using plateau selection of *L*1-norm in HAL.

#### Simulation

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Concluding Remarks Setting:

- $W \sim N(0, 4^2, -10, 10)$  drawn i.i.d. from a truncated normal distribution, bounded within [-10,10].
- A ~ Bernoulli(p(W)) with probability p(W) as a function of W bounded between [0.3, 0.7], given by p(W) = 0.3 + 0.1Wsin(0.1W) + ε, ε ~ N(0, 0.05<sup>2</sup>)
- $Y = 3sin(a_1W) + A + \varepsilon_2$  is a sinusoidal function of W, where  $\varepsilon_2 \sim N(0, 1)$ .
- *a*<sub>1</sub> controls the frequency (and true sectional variation norm) of the sinusoidal function.

The value of the parameter of interest, the ATE, is 1.

• The experiment is repeated 500 times, and interval coverages are computed

#### Simulation for n = 100

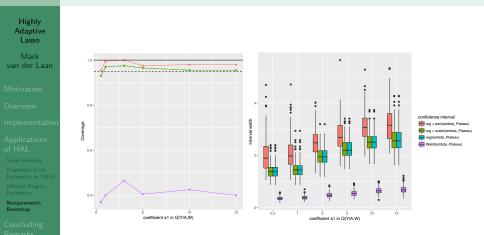


Figure: Coverage (left) and interval width (right) as a function of the  $a_1$  coefficient (i.e., sectional variation norm) of the true data-generating distribution.

### Concluding Remarks

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Concluding Remarks

- HAL is the first general nonparametric MLE.
- It converges at fast rate.
- HAL-TMLE is guaranteed asymptotically efficient.
- It provides finite sample robust TMLE for causal inference.
- It represents a class of HAL estimators by restricting function space, a priori, or data adaptively (screening).
- Its fit is a sparse representation, at most n-1 coefficients.
- It allows for nonparametric bootstrap.
- It has many applications, such as meta-HAL super-learner and outcome adaptive HAL-TMLE.

#### Frequently Asked Questions

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Concluding Remarks

- Why are global smoothing assumptions better than local smoothing assumptions?
- If I include HAL in my Super Learner library, then why would I include anything else?
- How does screening HAL's basis functions affect it's rate of convergence to the true function?
- How does HAL compare to BART?